Feature allocation models for tumor heterogeneity

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Tumor Heterogeneity 1

Slide 2

Tumor Heterogeneity

- Mutations acquired over a tumor's life history
- Every new mutation gives rise to a new subpopulation of cells ("subclone" = pair of haplotypes)
- \rightarrow heterogeneous population of cells, composed of subpopulations with varying numbers of mutations (e.g., Gerlinger et al. (2012, NEJM).
- Tumor history imprinted in each sample as the mosaicism of mutations.

Slide 3

Data

SNV: point mutations, $s = 1, \ldots, S$

Data: $N_{st} = \#$ reads mapped to locus of SNV s in sample t. $n_{st} = \#$ of these with SNV.

Sampling model:
$$n_{st} \sim Bin(N_{st}, p_{st})$$

Prior: in words.

- (i) p_{st} arises as a composition of sample t as a mixture of C latent haplotypes. (pairs of haplotypes define subclones).
- (ii) Mutation s in haplotype c is either present $(Z_{sc} = 1)$ or not $(Z_{sc} = 0)$. $\mathbf{Z}_c = (Z_{sc}, s = 1, \dots, S)$ defines haplotype c.
- (iii) Prior $p(\mathbf{Z})$ on $(S \times C)$ binary matrix \mathbf{Z} , prior $p(\boldsymbol{w})$ on mixture weights w_{tc} for composition (i).

Slide 4

Inference

Goal: Reconstruct cell subpopulations = estimate Z and C.

Problem: Deconvolution of p_{st} as a mixture of binary indicators Z_{sc}

$$p_{st} = \sum_{c} w_{tc} Z_{sc} + w_{t0} p_0 \tag{1}$$

plus "background noise"

Real problem: Z is latent, need to infer Z from the data.

Identifiability: In principle even feasible with one sample. Weights are identified across mutations s.

- Alternatives: • cluster variant allele fractions (VAF), $f_{st} = \frac{n_{st}}{N_{st}};$ \rightarrow subclones (e.g., PyClone: Roth et al, 2014 Nature Meth)
 - mixture of Beta's for observed VAF (SciClone: Miller et al., 2014 PLOS Comp Bio); variational Bayes
 - CNV data, fit mle for the prob of read alignments (THeta: Oesper et al., 2013 Genome Bio); mle mixture decomposition.
 - instead (1) explicitly models decomposition of samples into (hypothetical) subclones.

Feature allocation $\mathbf{2}$

Slide 5

Feature allocation

Feature allocation: binary matrix Zrows = mutations $s = 1, \ldots, S$; $cols = haplotypes (\subseteq \{1, \dots, S\})$

• Each mutation is in *multiple* subsets (haplotypes)

• Binary matrix $\mathbf{Z} = [Z_{sc}]$ records membership of mu-Composition of sample t as mix of haplotypes:

tations in haplotypes.

 $(w_{tc}, c = 1, \ldots, C) \sim \text{Dir}(\cdot),$ for each sample, $t = 1, \ldots, T.$

vs.

Clustering: each mutation is in *exactly one* subset (partition), row sum $\sum_{c} Z_{sc} = 1$

Slide 6

Indian buffet process

(Griffiths & Ghahramani, 2005 NIPS) Equivalent (original) definition of $p(\mathbf{Z})$. Let $Z_{sc} = \text{cus}$ tomer s selects dish c.

First customer: selects $C_1 \sim \text{Poi}(\alpha)$ dishes, $Z_{1c} = 1$, c = 4 $1, \ldots, C_1.$

Let $C = C_1$

s-th customer: Let
$$m_{sc} = \sum_{r < s} Z_{rc}$$
;

- $C_s \sim \operatorname{Poi}(\alpha/s)$ new dishes, $Z_{sc} = 1$,
- $c = C + 1, \ldots, C + C_s$
- Set $C \equiv C + C_s$

Prior for feature allocation Z.

Model for TH 3

Slide 7

Prior

Latent haplotypes: $p(\mathbf{Z})$ on $(S \times C)$ binary matrix, w. random C.

Feature allocation prior: Indian buffet process p(Z)(IBP),

with customers (experimental units) $s = 1, \ldots, S$ selecting dishes (features) $c = 1, \ldots, C$ Think of SNV s selecting haplotype (feature) c

IBP: define $p(\mathbf{Z})$, first for fixed C,

- $\pi_c \sim \operatorname{Be}\left(\frac{\alpha}{C}, 1\right)$ for each feature $c = 1, \ldots, C$
- $p(Z_{sc} = 1 \mid \pi_c) = \pi_c, s = 1, \dots, S$
- Drop unselected features

 $C \to \infty$ defines the IBC (Indian buffet process) $p(\mathbf{Z})$

Results - Pancreatic Cancer n = 5 samples of pancreatic cancer (PDAC, pancreatic ductal adenocarcinoma).

see figures on slide 1.

Adding copy number variation

Slide 9

Haplotypes vs. Subclones

So far our discussion is on haplotypes. • select dish c = 1, ..., C with prob $p(Z_{sc} = 1) = \frac{m_{sc}}{s}$; But cell subpopulations (subclones) are defined by pairs of haplotypes.



a diploid organism, pairs of haplotypes define a unique genome.

- Next: will change to subclones (pairs of haplotypes) as experimental units.
- Subclone can have $Z_{sc} \in \{0, 1, 2\}$ copies of each mutation.

Slide 10

Alternative models and generalizations

To represent subclones and more we relax assumptions, using

(i) cIBP for **subclones** (= pairs of haplotypes)

- (ii) CNV's: use data on copy number variation N_{st}
- (iii) repulsive priors (DPP): IBP includes independence $\overline{Slide 14}$ across columns!



DPP on images "Jaguar" $X = \{x_1, \dots, x_K\}, x_k \in$ "Jaguar images" Want more diversity, fewer cats – more football teams etc. :-)



DPP: Estimated \hat{Z}

IBP: Estimated \hat{Z}

Slide 17

DPP

Slide 20

Summary

proaches remain feasible.

netic tree of subclones

Limitations: and extensions

 $- \operatorname{arrgh!})$

DP mixture: n clusters ble – and seems to work.

- **DPP:** point process $X = \{x_1, \ldots, x_K\}$ on $x_i \in S$ for some space S, e.g. $S \subseteq \Re^2$ or S = images.
- **Idea:** instead of many similar points x_j , generate only few distinct ones.



- **Repulsive point process:** avoid duplication of similar values x_j ;
 - for example, google-ing "Jaguar", you want $X = \{ cat, car, football team, ... \}$

Slide 18

Truth Z^o

DPP

DPP: point process $X = \{x_1, \ldots, x_K\}$ on $x_i \in S$ for some space S;

penalizes "similar" x_j .

DPP on finite discrete space: $p(X) \propto \det C_X$ where $C_{X,ij} = C(x_i, x_j)$ for a p.d. kernel C(x, x').

DPP on (bounded) continuous space: density w.r.t. unit rate Poisson process

$$f(X) = \det C_X / \prod (1 + \lambda_h)$$

with λ_h = eigenvalues of operator $T : h \rightarrow \int_S C(x,y)h(y)dy$.

MCMC: easy for finite S (e.g., Kulesza, A. and Taskar, B. (2012 Machine Learn.) for continuous S: reversible jump MCMC with f(X) (Xu et al., 2015 arXiv)

Summary **TH:** Model-based estimation of cell subpopulations is possi-

Big data: MCMC is not feasible anymore - alternative ap-

Tumor phylogenetics: Without condition on phyloge-

A priori independent cell types: indpendent $z_c = (Z_{1...S,c})$, with $p(z_c = z_{c'}) > 0$, a priori (i know

Alternative dependent prior using DPP or others.

DPP - ExamplesSimulation truth = mix of normals (left). Clustering model Slide 21
with DPP (center) vs. DP mixture (right) prior.



Density estimation for a mix of normals.

Slide 19

Simulation truth Z^{o} (left). Feature allocation with DPP prior vs. IBP on features Slide 24

| 6 Extra Slides – MAD Baves | Affandi, R. H., Fox, E., and Taskar, B. (2013). Approximate inference in continuous determinantal processes. Adv in Neural Inf Processing Systems, 1430–1438. |
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| Slide 22 | Lavancier, F., Møller, J., and Rubak, E. (2015). Determinantal point process models and statistical in- ference. JRSSB, 77, 853–877. |
| Posterior inference for IBP using MAD Bayes with YANXUN XU, UT Austin; YUAN YUAN, Baylor C.of Med.; YUAN JI and KAMALAKAR GULUKOTA, NorthShore Hospi- tal. | Lee, J., Müller, P., Ji, Y. and Gulukota, K. (2015) "A Bayesian Feature Allocation Model for TH," Ann. of Applied Stat, 9, 621-639. |
| | ^{bf}Lee, J., Müller, P., Sengupta, S., Gulukota, K. and Ji, Y. (2014). "Bayesian Inference for Intra-Tumor Heterogeneity in Mutations and Copy Number Variation," <i>JRSSC</i>, final(?) revision. |
| DP mixture: Kulis & Jordan (2012) recognize log posterior ≈ criterion function in k-means – voila! This is for normal sampling, asymptotically for small variance and shrinking total mass. IBP: Broderick et al. (2013) extend a similar argument the IBP, with normal sampling and small variance and shrinking rate of new features, | Roth, A. et al. (2014). "PyClone", Nature Methods, 11, 396- Sengupta, S., Guluokta, K., Lee, J., Müller, P., and Ji, Y. (2015) "Bayclone: Bayesian Nonparametric Inference of Tumor Subclones Using NGS Data." In Proceedings of The Pacific Symposium on Biocomputing (PSB) 2015, 467-78. ^OXu Y, Müller P, Yuan Y, Gulukota K and Ji Y, (2015). "MAD Bayes for Tumor Heterogeneity," JASA, 110, 503-514. |
| IBP with binomial sampling: same argument can b made :-) using increasing scaling of Bin with β and shrinking IBI | e P |

par γ , using $\gamma = \exp(-\beta\lambda^2)$

Approx posterior: use k-means with different starting values to characterize posterior.

Slide 23

Results – Pancreatic Cancer

n=5 samples of pancreatic cancer (PDAC, pancreatic ductal adenocarcinoma). Estimated $w_{tc}{:}$

